22S:105 Statistical Methods and Computing

Contingency Tables and the Chi-Square Test Introduction to ANOVA

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- death certificate incorrect and required recoding of underlying cause of death
- Question of interest: Are there differences between the two hospitals with respect to practices in completing death certificates
- One way to address the question: Test null hypothesis that, within each category of death certificate status, the proportions of death certificates coming from Hospital A are the same.

The Chi-square test for differences among more than 2 proportions

We are interested in the *independent samples* case.

Example:

- A study investigated the accuracy of death certificates by comparing the results of 575 autopies to the causes of death listed on the certificates.
- Two hospitals participated in the study.
 - community hospital, labeled A
 - university ospital, labeled B
- Three possible cases
 - death certificate confirmed accurate
 - death certificate contained inaccuracies but did not require correction of underlying cause of death

Another multiple comparisons problem!

 $H_0: p_c = p_i = p_r$ $H_a: p_c \neq p_i \text{ or } p_c \neq p_r \text{ or } p_i \neq p_r$

- We will *first* test whether there are *any* significant differences.
- Only if we reject H_0 in the overall test will we do pairwise tests to find out *which* population proportions are different.

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Results

	Hospital A	Hospital B	Total
Confirmed accurate	157	268	425
Inacc, no recoding	18	44	62
Incorrect, recoding	54	34	88
Total	229	346	575

The overall sample proportion of death certificates from hospital A is

$$\frac{229}{575} = 0.398$$

If H_0 is true, we would expect this same proportion of hospital A certificates in all three categories.

According to Table E, the .05 cutoff under a Chi-square distribution with 2 d.f. is 5.99.

We can reject H_0 because 21.62 > 5.99. The p-value < 0.001.

We conclude that the proportions of death certificates from Hospital A are not the same for the three different categories of certificate status.

Observed and expected counts

	Hospital A	Hospital B	Hospital A	Hospital 1
Accurate	157	268	169.3	255.7
Incorrect	18	44	24.7	37.3
Recode	54	34	35.0	53.0

The Chi-square statistic is

$$X^2 = 21.62$$

- r = 3 rows
- c = 2 columns
- So the degrees of freedom is (r-1)(c-1) = 2(1) = 2

This Chi-square test in SAS

```
options linesize = 72;
data dthcert ;
input hosp $ status $ count ;
datalines;
A C 157
A I 18
A R 54
B C 268
B I 44
B R 34
proc freq data = dthcert ;
tables status * hosp / expected;
weight count ;
run ;
proc freq data = dthcert ;
tables status * hosp / chisq ;
weight count ;
run ;
```

9 10

229

39.83

STATISTICS FOR TABLE OF STATUS BY HOSP

0.193

346

60.17 100.00

575

TABLE OF STATUS BY HOSP Total

STATUS HOSP

Frequency Expected | Percent | Row Pct | |B | Total Col Pct |A -----+ C | 157 | 268 | 425 | 169.26 | 255.74 | | 27.30 | 46.61 | 73.91 | 36.94 | 63.06 | | 68.56 | 77.46 | -----+ | 18 | 44 | | 24.692 | 37.308 | | 3.13 | 7.65 | 10.78 | 29.03 | 70.97 | | 7.86 | 12.72 | ----+ | 54 | 34 | | 35.047 | 52.953 | | 9.39 | 5.91 | 15.30 | 61.36 | 38.64 | | 23.58 | 9.83 |

TABLE OF STATUS BY HOSP

-----+

Sample Size = 575

Cramer's V

STATUS HOSP Frequency Percent | Row Pct | Col Pct |A |B | Total -----C | 157 | 268 | 425 | 27.30 | 46.61 | 73.91 | 36.94 | 63.06 | | 68.56 | 77.46 | -----+ | 18 | 44 | | 3.13 | 7.65 | 10.78 | 29.03 | 70.97 | | 7.86 | 12.72 | -----| 54 | 34 | | 9.39 | 5.91 | 15.30 | 61.36 | 38.64 | | 23.58 | 9.83 | Total 229 346 575 39.83 60.17 100.00

The sample proportions are

	Hospital A	Hospital B	Total
Confirmed accurate	157	268	0.369
Inacc, no recoding	18	44	0.409
Incorrect, recoding	54	34	0.614
Total	229	346	575

More advanced methods provide tests and confidence intervals to formalize analysis of which population proportions are significantly different.

Goal: to compare population means under three different "treatments"

- a *three*-independent-sample problem
- Call the population mean heart rates μ_1 for when pets are present, μ_2 for when friends are present, and μ_3 for when women perform task alone: then

 $-H_0: \mu_1 = \mu_2 = \mu_3$

 $-H_a: \mu_1 \neq \mu_2 \text{ or } \mu_1 \neq \mu_3 \text{ or } \mu_2 \neq \mu_3$

* not one-sided or 2-sided

Comparing more than two population means

Example: Does the presence of pets or friends affect responses to stress?

- Allen, Blascovich, Tomaka, and Kelsey, 1988, Journal of Personality and Social Psychol-ogy
- subjects: 45 women who described themselves as dog lovers
- randomly assigned to three groups: to do a stressful task
 - 1. alone
 - 2. with a good friend present
 - 3. with their dog present
- Subjects' mean heart rate during the task was one measure of the effect of stress.

SAS descriptive statistics:

Analysis Variable : BEATS

	GRC				
	Mean				
	82.5240667	9.2415747	62.6460000	99.0460000	
GROUP=F					
N	Mean		Minimum		
15	91.3251333	8.3411341	76.9080000	102.1540000	
GROUP=P					
N	Mean	Std Dev	Minimum	Maximum	
 15	73.4830667	9.9698202	58.6920000	97.5380000	

To infer about the three population means, we might use the two-independent-sample t test 3 times:

- Test $H_0: \mu_1 = \mu_2$ to see if mean heart rate when pet is present differs from mean when friend is present.
- Test $H_0: \mu_1 = \mu_3$ to see if mean heart rate when pet is present differs from mean when alone.
- Test H_0 : $\mu_2 = \mu_3$ to see if mean heart rate when friend is present differs from mean when alone.

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Multiple comparisons procedures in statistics

- issue: how to do many comparisons at once with some overall measure of confidence in all our conclusions
- two steps
 - overall test of whether there is good evidence of *any* differences among parameters we wish to compare
 - follow-up analysis to decide which of parameters differ and to estimate size of differences

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Problem with this approach:

- 3 p-values for 3 different tests don't tell us how likely it is that *three* sample means are spread apart as far as these are.
- might be that $\bar{x}_1 = 73.48$ and $\bar{x}_2 = 91.32$ are significantly different if we look at just 2 groups but *not* significantly different if we know they are the smallest and largest means in 3 groups
 - As more and more groups are considered, we expect gap between smallest and largest sample mean to get larger.
 - (Imagine comparing heights of shortest and tallest person in larger and larger groups of people.)
- the probability of Type I error for the whole set of t-tests will be much bigger than the α level set for each one

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Step one: One-Way Analysis of Variance (ANOVA)

- step one (overall test) for *some* difference among 3 or more population means
- \bullet uses an F test to compute a p-value

Analysis of Variance Procedure

Class Levels Values
GROUP 3 C F P

Number of observations in data set = 45

Analysis of Variance Procedure

Dependent	Variable	: BEATS				
		Sı	ım of	Mean		
Source	Di	F Sqi	iares	Square F	7 Value	Pr > F
Model		2 2387.688	20020 11	.93.8444960	14.08	0.0001
Error	4			84.7928450	14.00	0.0001
	-			04.7320430		
Corrected	Total 4	1 5948.988	34836			
	R-Square	Э	C.V.	Root MSE	BEAT	TS Mean
	0.40136	11.:	16915	9.2083030	82.	444089
Source	Di	F Anos	ra SS M	lean Square F	Value	Pr > F
GROUP	:	2 2387.688	39920 11	93.8444960	14.08	0.0001

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F distributions

- many different F distributions, identified by two parameters
 - numerator degrees of freedom = I 1
 - denominator degrees of freedom = N I

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Main idea of ANOVA

What matters is how far apart sample means are relative to variability of individual observations.

• F statistic

 $F = \frac{variation \; among \; the \; sample \; means}{variation \; among \; individuals \; in \; the \; same \; sample}$

• compare to a cutoff value in an **F** distribution

Notation:

- I = number of different populations whose means we are studying
- n_i = number of observations in sample from ith population
- N = total number of observations in all samples combined

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Example

Do four varieties of tomato plant differ in mean yield? Agronomists grew 10 plants of each variety and recorded the yield of each plant in pounds of tomatoes.

What are

- the populations of interest
- the variable of interest
- I
- each n_i
- the degrees of freedom for the ANOVA F statistic

Assumptions for One-Way ANOVA

- We have I independent simple random samples, one from each of I populations.
- Each population i has a normal distribution with unknown mean μ_i .
 - As with t-tests, if sample sizes are large enough in each sample, Central Limit Theorem says inference based on sample means is OK even if population distributions are not exactly normal.

Step two: individual t-tests with correction for multiple comparisons

This is the follow-up test.

• should be carried out *only* if the F test from one-way ANOVA is significant at the chosen significance leve.

Goal: to set the *overall* probability of committing a type I error at α when doing pairwise comparisons of k different means

- \bullet we will perform $\left(k\atop 2\right)$ two-independent-sample t-tests
- we will conduct each one at the significance level

 $\alpha^* = \frac{\alpha}{\binom{k}{2}}$

• This is called the Bonferroni correction

• All of the populations have the same standard deviation σ (unknown)

- unlike t-tests, there is no general procedure when population standard deviations are not assumed to be equal
- rough rule of thumb: if largest sample standard deviation is no more than twice the smallest sample standard deviation, then population standard deviations probably are close enough to equal that ANOVA procedure is OK

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- very conservative

- There are k=3 samples, so there are $\binom{k}{2}=3$ different pairs to compare.
- To get an overall significance level $\alpha = .05$ on all 3 tests considered together, we conduct each one at

$$\alpha^* = \frac{.05}{3} = .0167$$

- That is, we would consider the difference between two population means to be significantly different from zero at the .05 level only if the p-value for the t-test for that pair was less than .0167. Equivalently, we could multiply the p-value from each t-test by 3.

* If the result was less than .05, we would consider the difference between two population means to be significantly different from zero at the .05 level

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SAS does the adjusting and prints a grouped list of the classes. Means with the same letter are not significantly different at the specified alpha level.

Analysis of Variance Procedure

Bonferroni (Dunn) T tests for variable: BEATS

NOTE: This test controls the type I experimentwise error rat generally has a higher type II error rate than REGWQ.

Alpha= 0.05 df= 42 MSE= 84.79285 Critical Value of T= 2.49 Minimum Significant Difference= 8.3847

Means with the same letter are not significantly different.

Bon Grouping	Mean	N	GROUP
A	91.325	15	F
В	82.524	15	С
C	73.483	15	P

One-way ANOVA in SAS

```
options linesize = 79;
data pet ;
infile '/temp/pet.dat';
input group $ beats ;
run :
proc sort data = pet ;
by group;
run ;
proc means data = pet ;
by group ;
var beats ;
run ;
proc anova data = pet ;
class group ;
model beats = group ;
proc anova data = pet ;
class group ;
model beats = group ;
means group / bon alpha = .05 ;
run ;
```